

## **Time series models: Continuous data**

Atmospheric variables tend to be persistent, they have a lag-autocorrelation  $r_1 > 0$ .

Simplest time series model:  $x_{t+1} - \mu = \phi_1(x_t - \mu) + \varepsilon_{t+1}$ , or in terms of the anomalies,  $x'_{t+1} = \phi_1 x'_t + \varepsilon_{t+1}$

This is an **autoregressive model of order 1 (AR(1))**. Such model can be used to:

- Fit the time series and derive some of its properties. Similar to fitting a theoretical probability distribution to a sample.
- To make a forecast:  $\hat{x}_{t+1} - \mu = \phi_1(x_t - \mu)$

We need to determine  $\phi_1 = \phi$  and the variance of the error  $\text{var}(\varepsilon)$  from the data that we want to fit with the AR(1) model.

Since  $r_1 = \text{corr}(x'_t, x'_{t+1}) = \frac{\overline{(x'_t x'_{t+1})}}{\sqrt{\sigma_{x'_t}^2 \sigma_{x'_{t+1}}^2}} = \frac{\overline{(x'_t x'_{t+1})}}{\sigma_{x'_t}^2}$ , a linear regression forecast

is simply  $\hat{x}_{t+1} - \mu = r_1(x_t - \mu)$  or  $\phi_1 = r_1$

Note that autocorrelations at longer lags are not zero for AR(1), even though we only need the last observation to make an AR(1) forecast:

$$P\{X_{t+1} \leq x_{t+1} | X_t \leq x_t, X_{t-1} \leq x_{t-1}, \dots, 1\} = P\{X_{t+1} \leq x_{t+1} | X_t \leq x_t\}$$

i.e., just the last observation is enough to make a forecast, but

$$r_1 = \phi_1, \quad r_2 = \phi_1^2, \quad r_3 = \phi_1^3, \dots$$

In the AR(1)

$$x'_{t+1} = \phi x'_t + \varepsilon_{t+1}.$$

Multiply this equation by  $x'_t$  and average over a long time series, divide by  $\sigma_x^2 \approx s_x^2$ , and use  $\overline{x'_t x'_t} = \sigma_x^2 = \overline{x'_{t+1} x'_{t+1}}$ ,  $\overline{x'_t x'_{t+1}} = r_1 \sigma_x^2 = \phi \sigma_x^2$  and obtain

$\overline{x'_t \varepsilon_{t+1}} = 0$ , i.e., the error is uncorrelated with the predictor.

Multiply  $x'_{t+1} = \phi x'_t + \varepsilon_{t+1}$  now by  $x'_{t+1}$  and average over a long time series and obtain

$\overline{x'_{t+1} x'_{t+1}} = \overline{\phi x'_{t+1} x'_t} + \overline{\varepsilon_{t+1}^2}$  so that the unexplained variance of the prediction is

$$\sigma_{\varepsilon}^2 = (1 - \phi^2) \sigma_x^2.$$

The estimate of  $\hat{\phi} = r_1$  obtained from a sample can be tested for significance (whether it is REALLY different from zero) as in linear regression. Recall that in linear regression  $y_t - \bar{y} = b_1(x_t - \bar{x}) + \varepsilon_t$ , and the variance of  $b_1$  is

estimated as  $\sigma_{b_1}^2 \approx \frac{\overline{\varepsilon_t^2}}{\sum_{t=1}^n (x_t - \bar{x})^2} \approx \frac{\overline{\varepsilon_t^2}}{n\sigma_x^2}$  where  $\overline{\varepsilon_t^2} \approx \frac{1}{n} \sum_{t=1}^n (y_t - \hat{y}_t)^2$  is the

forecast error squared (unexplained variance). Then we use a t-test

$T = \frac{b_1 - 0}{\sigma_{b_1}}$ . (Note that this is the same type of statistics that we would use to

estimate the significance of a climate trend  $b_1$  if we assume that the trend is linear with time,  $Temp = Temp_{mean} + b_1 \text{ time}$ .)

As shown above, for AR(1)  $x'_{t+1} = \phi x'_t + \varepsilon_{t+1}$ , the variance “unexplained” by regression is  $\sigma_{\varepsilon}^2 = (1 - \phi^2) \sigma_x^2$ .

Using the same formula for the estimated variance of the linear coefficient  $\phi$  as for linear regression coefficient  $b_1$ , the error variance of  $\phi$  is estimated as

$$\sigma_{\phi}^2 = \frac{\sigma_{\varepsilon}^2}{n\sigma_x^2} = \frac{1 - \phi^2}{n}.$$

One can then test whether the persistence is significantly different from zero

using  $z = \frac{\hat{\phi} - 0}{\sqrt{(1 - \hat{\phi}^2) / n}}$  (we can use a Gaussian distribution because  $\hat{\phi}$

determines both the mean and its standard deviation).

The variance of the noise (unexplained variance), when we use

AR(1):  $\sigma_\varepsilon^2 = (1 - \phi^2)\sigma_x^2$  for the population, and for a sample,

$$s_\varepsilon^2 = \frac{(1 - \hat{\phi}^2)}{n - 2} \sum_{t=1}^n (x_t - \bar{x})^2 = \frac{n - 1}{n - 2} (1 - \hat{\phi}^2) s_x^2 \text{ but usually } n \text{ is large so } \frac{n - 1}{n - 2} \approx 1$$

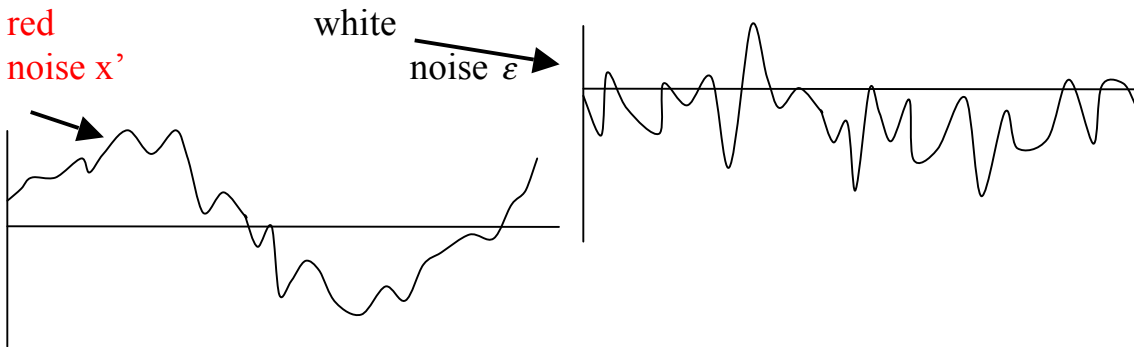
### Applications:

- a) Create a persistent time series (red noise) that looks like nature:

$$x'_{t+1} = \hat{\phi}x'_t + \varepsilon_{t+1}$$

red  
noise  $x'$

white  
noise  $\varepsilon$



- b) Make a forecast:  $x'_{t+1} = \hat{\phi}x'_t$

(How would blue noise look like? Hint: for white noise, persistence,  $r_1=0$ ; for red noise,  $r_1>0$ ; for blue noise,  $r_1<0$ , the anomaly changes sign very frequently)

## Higher order autoregressive model AR(K)

$$x_{t+1} - \mu = \sum_{k=1}^K \phi_k (x_{t-k+1} - \mu) + \varepsilon_{t+1}$$

For example, a 2<sup>nd</sup> order autoregressive model, AR(2):

$$x_{t+1} - \mu = \phi_1 (x_t - \mu) + \phi_2 (x_{t-1} - \mu) + \varepsilon_{t+1}$$

$$\text{or } x'_{t+1} = \phi_1 x'_t + \phi_2 x'_{t-1} + \varepsilon_{t+1} \quad (1)$$

Multiply both sides of (1) by  $x'_t$ , take an average over a long time series, divide by  $s_x^2$ , and obtain the following relationship:

$$r_1 = \hat{\phi}_1 + r_1 \hat{\phi}_2 \quad (2)$$

Similarly, multiply both sides of (1) by  $x'_{t-1}$ , take an average over a long series, divide by  $s_x^2$ , and obtain the following:

$$r_2 = r_1 \hat{\phi}_1 + \hat{\phi}_2 \quad (3)$$

From (2) and (3) can solve for  $\hat{\phi}_1, \hat{\phi}_2$ :

$$\hat{\phi}_1 = \frac{r_1(1-r_2)}{1-r_1^2}, \quad \hat{\phi}_2 = \frac{r_2-r_1^2}{1-r_1^2}$$

Or, if we know  $\hat{\phi}_1, \hat{\phi}_2$

$$r_1 = \frac{\hat{\phi}_1}{1-\hat{\phi}_2}, \quad r_2 = \hat{\phi}_2 + \frac{\hat{\phi}_2^2}{1-\hat{\phi}_2}$$

The expected variance of the error is

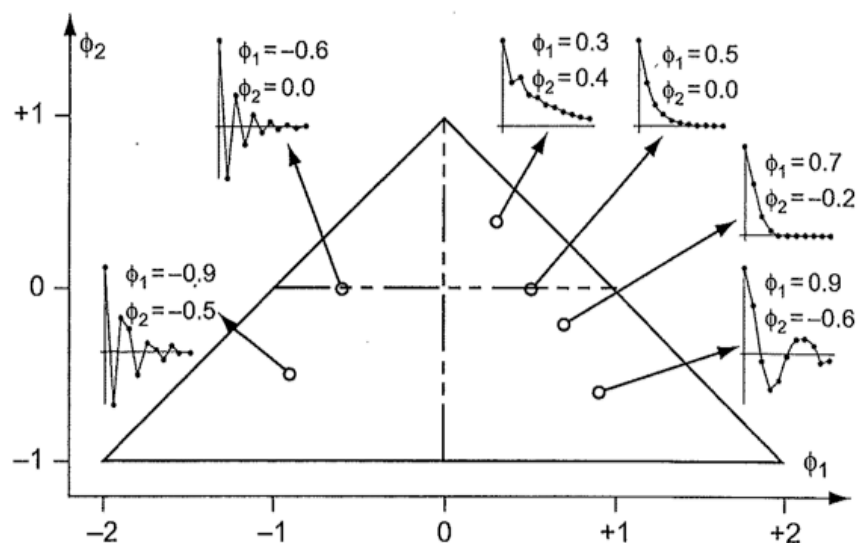
$$\sigma_\varepsilon^2 = (1 - \hat{\phi}_1 r_1 - \hat{\phi}_2 r_2) \sigma_x^2$$

For an AR(2) series to be stationary (so that it does not drift away), the following conditions have to be satisfied:

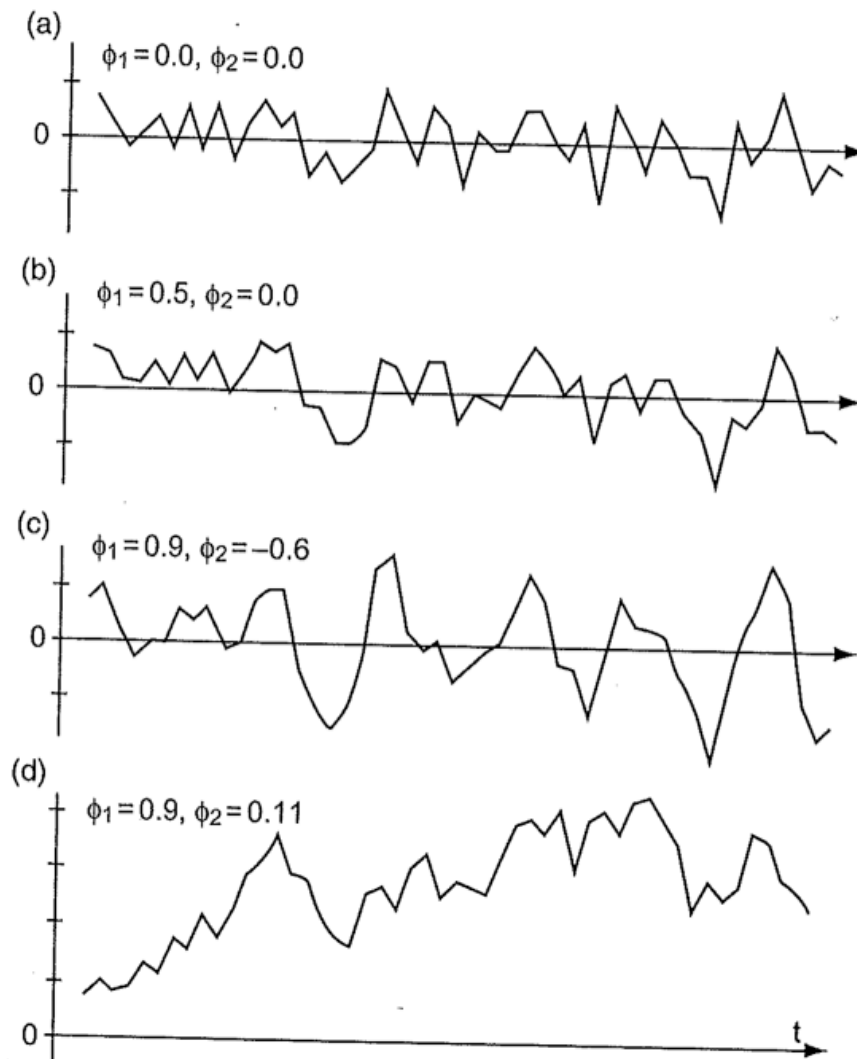
$$-1 \leq \phi_2 \leq 1, \quad \phi_1 + \phi_2 \leq 1, \quad \phi_2 - \phi_1 \leq 1$$

(see figures 9.7, 9.8 in Wilks).

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**9.7** The allowable parameter space for stationary AR(2) processes, with insets showing autocorrelation functions for AR(2) models. The horizontal  $\phi_2 = 0$  line locates the AR(1) models as special cases, and autocorrelation functions for these are shown. AR(2) models appropriate to atmospheric time series usually exhibit  $\phi_1 > 0$ .

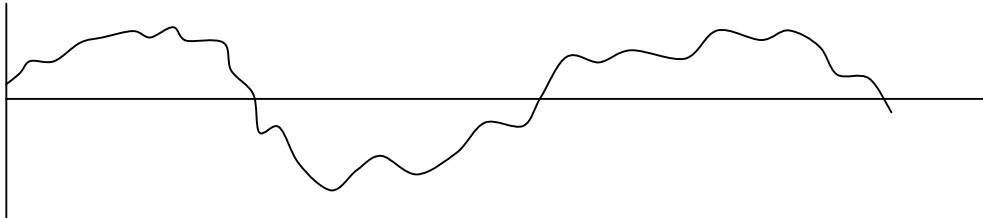


### Variance of a time series

If a time series has zero autocorrelation, the variance of the mean is the familiar

$\sigma_{\bar{x}}^2 = \frac{s_x^2}{n}$  which shows that the mean of the time series measured by  $n$  time steps has a variance  $n$  times smaller than the individual measurements.

However, if  $r_1 > 0$  then the time mean has a larger variance than indicated by this formula! This is because there are fewer independent measurements than  $n$ :



This effect can be estimated by a variance inflation factor  $V$  :

$s_{\bar{x}}^2 = V \frac{s_x^2}{n}$  where  $V = \frac{1 + \phi_1}{1 - \phi_1}$  for AR(1). In other words, the effective number of independent observations in an AR(1) time series is

$$n' = \frac{1 - \phi_1}{1 + \phi_1} n$$

This is very important when estimating the number of degrees of freedom for, say, daily observations. If  $\phi_1 = 0.5$ , then  $n' \approx \frac{0.5}{1.5} n = n/3$ . These means that we should consider only observations every third day, or conversely, assume that the number of degrees of freedom is  $n/3$ .

Note on the “physical meaning” of autoregression modeling  
(don’t know the author but it is a nice summary)

### Autoregression Modeling:

● mathematical model used to try and explain a time series of observations:  $\{y_{i\Delta t}\}$ ,  $i=1,2,\dots,N$ .

AR(p) process:

$$y_t = a_1 y_{t-\Delta t} + a_2 y_{t-2\Delta t} + \dots + a_p y_{t-p\Delta t} + a_0 v_t$$

↑  
Gaussian white noise

● an AR(p) model is a discretized  $p^{\text{th}}$ -order ordinary differential equation:

e.g. AR(1):  $\frac{d}{dt}y(t) + \frac{y(t)}{\tau} = v(t)$  ← Gaussian white noise

$$y(t) = y_0 \exp\left(-\frac{t}{\tau}\right)$$

$\tau$  = characteristic timescale, or ‘memory’  
~ (heat capacity) / (cooling rate)

AR(1), aka ‘red noise’ = simplest self-consistent model for a geophysical system (e.g., Hasselman, 1976).

AR(2+): combination of oscillations/growing and decaying exponentials.



### Autoregressive, moving-average models ARMA(K,M)

In these models we assume that the noise has some persistence, and persist the last few observed “noises”:

$$x_{t+1} - \mu = \sum_{k=1}^K \phi_k (x_{t-k+1} - \mu) + \varepsilon_{t+1} + \sum_{m=1}^M \theta_m \varepsilon_{t-m+1}$$

The simplest ARMA model is ARMA(1,1):

$$x_{t+1} - \mu = \phi_1 (x_t - \mu) + \varepsilon_{t+1} + \theta_1 \varepsilon_t$$

The lag-1 autocorrelation for this model is

$$r_1 = \frac{(1 - \phi_1 \theta_1)(\phi_1 - \theta_1)}{1 + \theta_1^2 - 2\phi_1 \theta_1},$$

and the expected error variance is given by

$$\sigma_\varepsilon^2 = \frac{1 - \phi_1^2}{1 + \theta_1^2 + 2\phi_1 \theta_1} \sigma_x^2.$$

So, in practice, to make an ARMA(1,1) forecast, we need to fit the data and compute the first two lag autocorrelations  $r_1$ ,  $r_2$ . Then, obtain  $\phi_1$  from

$$r_2 = \phi_1 r_1, \text{ and finally } \theta_1 \text{ from } r_1 = \frac{(1 - \phi_1 \theta_1)(\phi_1 - \theta_1)}{1 + \theta_1^2 - 2\phi_1 \theta_1}.$$

Then the forecast for future data becomes:

$$\hat{x}_{t+1} - \mu = \phi_1 (x_t - \mu) + \theta_1 \varepsilon_t$$

where  $\varepsilon_t$  is the last observed forecast error  $\varepsilon_t = x_t - \hat{x}_t$ .